

Speeds of coming down from infinity for continuous-state nonlinear branching processes

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Outline of the Talk

1 Introduction

- Continuous-state branching process
- Nonlinear continuous-state branching process
- Elements of SNLP

2 Main Results

- Moments of the hitting time for nonlinear CSBP
- Laplace transform of the hitting time
- Speeds of coming down from infinity

Feller branching process and Branching property

- A Feller branching process is the unique nonnegative solution to SDE

$$X_t = X_0 + \int_0^t \sqrt{\gamma X_s} dB_s^X,$$

where B^X is a Brownian motion and $\gamma > 0$ is the branching rate.

- Let Y be another independent Feller branching processes.

$$Y_t = Y_0 + \int_0^t \sqrt{\gamma Y_s} dB_s^Y$$

Since $X + Y$ is a continuous martingale with $\langle X + Y \rangle_t = \langle X \rangle_t + \langle Y \rangle_t = \int_0^t \gamma(X_s + Y_s) ds$, then $X + Y$ solves SDE $Z_t = X_0 + Y_0 + \int_0^t \sqrt{\gamma Z_s} dB_s^Z$.

- Such an additive property is called **branching property**.

Continuous-state branching process (CB processes)

- Feller diffusion is an example of **continuous-state branching processes**. In general, a continuous-state branching process is a nonnegative Markov process X (with no negative jumps) satisfying the branching property.
- Its Laplace transform is determined by

$$\mathbb{E}_x e^{-\theta X_t} = e^{-x u_t(\theta)}$$

where function $u_t(\theta)$ satisfies the differential equation

$$\frac{\partial u_t(\theta)}{\partial t} + \psi(u_t(\theta)) = 0$$

with $u_0(\theta) = \theta$ and

$$\psi(\lambda) = b\lambda + \frac{1}{2}\sigma^2\lambda^2 + \int_0^\infty (e^{-\lambda x} - 1 + \lambda x)\Pi(dx)$$

for $\sigma > 0$, $b \in \mathbb{R}$ and for σ -finite measure Π on $(0, \infty)$

satisfying $\int_0^\infty (z \wedge z^2)\Pi(dz) < \infty$

Lamperti transform

- CB process is associated with a spectrally positive Lévy process via the **Lamperti** time change.
- Write Z for a spectrally positive Lévy process and $\tau_0^- := \inf\{t : Z_t = 0\}$. for its first time of reaching 0. Write

$$\eta(t) := \int_0^{t \wedge \tau_0^-} \frac{1}{Z_s} ds \quad \text{and} \quad \eta^{-1}(t) := \inf\{s \geq 0 : \eta(s) > t\}.$$

Then process $X_t := Z_{\eta^{-1}(t) \wedge \tau_0^-}$ is the CB process.

A class of nonlinear continuous-state branching processes

- We can generalize the above mentioned Lamperti transform.
Let R be a positive function on $[0, \infty)$. Define

$$\eta(t) := \int_0^{t \wedge \tau_0^-} \frac{1}{R(Z_s)} ds \quad \text{and} \quad \eta^{-1}(t) := \inf\{s \geq 0 : \eta(s) > t\}.$$

Then $X_t := Z_{\eta^{-1}(t) \wedge \tau_0^-}$ is a **nonlinear continuous-state branching process** with branching rate function $R(x)$.

- Process X has a generator L on $C^2[0, \infty)$ such that

$$\begin{aligned} Lf(x) &:= R(x)L^*f(x) \\ &:= R(x) \left(-bf'(x) + \frac{\sigma^2}{2}f''(x) + \int_0^\infty (f(x+u) - f(x) - uf'(x)) \Pi \right) \end{aligned}$$

where L^* is the generator for a spectrally positive Lévy process.

- Intuitively, it is a branching process whose branching rate depends on the current population size.
- Observe that for $b > 0$, $\tau_b^- := \inf\{t : Z_t < b\}$ and $T_b^- := \inf\{t : X_t < b\}$,

$$T_b^- = \int_0^{\tau_b^-} 1/R(Z_s) ds.$$

We need to understand the weighted occupation time for Z .

- Such a nonlinear continuous-state branching process had been studied in P. Li (2016).

- A spectrally positive Lévy process (SPLP) is a Lévy process with no negative jumps.
- For spectrally positive Lévy process Z_t with $Z_0 = 0$,

$$\mathbb{E}e^{-\theta Z_t} = e^{t\psi(\theta)},$$

for $\theta, t \geq 0$, where the **Laplace exponent** for $-Z$ is

$$\psi(\theta) = b\theta + \frac{1}{2}\sigma^2\theta^2 + \int_{-\infty}^0 (e^{\theta z} - 1 - \theta z) \Pi(dz),$$

for $\gamma \in \mathbb{R}$ and $\sigma \geq 0$. Also, the Lévy measure Π is a σ -finite measure on $(-\infty, 0)$ such that

$$\int_{-\infty}^0 (z \wedge z^2) \Pi(dz) < \infty.$$

- $\mathbb{E}Z_1 = -\psi'(0+)$.

Scale function

- The Laplace exponent ψ is strictly convex and $\lim_{\theta \rightarrow \infty} \psi(\theta) = \infty$. Thus, there exists an inverse function $\Phi: [0, \infty) \rightarrow [0, \infty)$ such that

$$\psi(\Phi(\theta)) = \theta, \quad \theta \geq 0.$$

- We often need **scale functions** to study the fluctuations of SNLP.
- The scale function W of the process $-Z$ is defined as the function with Laplace transform on $[0, \infty)$ given by

$$\int_0^{\infty} e^{-\lambda z} W(z) dz = \frac{1}{\psi(\lambda)} \quad \text{for } \lambda > \Phi(0).$$

Coming down from infinity

- We say the process X comes down from infinity if

$$\lim_{b \rightarrow \infty} \lim_{x \rightarrow \infty} \mathbb{P}_x(T_b^- < \infty) = 1. \quad (1)$$

- Let $\mathbb{P}_\infty := \lim_{x \rightarrow \infty} \mathbb{P}_x$ and \mathbb{E}_∞ be the corresponding expectation.
- We can show that X comes down from infinity if and only if $\mathbb{E}_\infty T_b^- < \infty$ for all large b .
- If X comes down from infinity, then ∞ can be treated as an entrance boundary. \mathbb{P}_∞ is the entrance law.
- The speed of coming down from infinity has been studied for Λ -coalescent in Berestycki et al (2010) and for birth and death processes in Bansaye et al (2015).

Why coming down from infinity?

- If process Z has a negative drift, it can go down by a large amount over a long period.
- Starting at x with a large value of $R(x)$, due to the time change, a short time behavior of X near time 0 corresponds to a long term behavior of Z , and we see a significant drop for X over the short time period;
- If X is subcritical and $R(x)$ increases fast enough as $x \rightarrow \infty$, the drop becomes so drastic and coming down from infinity occurs;
- If X is critical, coming down from infinity can still occur if $R(x)$ increases even faster as $x \rightarrow \infty$.

We are interested in two kinds of rate function $R(x)$.

- (slow regime) $R(x) = x^\theta$, $x, \theta > 0$.
- (fast regime) $R(x) = e^{\theta x}$, $x, \theta > 0$.

Outline of the strategy

Assume that X comes down from infinity.

- Under \mathbb{P}_∞ we first estimate T_b^- as $b \rightarrow \infty$.
- Let $\underline{X}_t := \inf_{0 \leq s \leq t} X_s$ be the running minimum process of X . Since $\underline{X}_{T_b^-} = b$, we can show that $\lim_{t \rightarrow 0+} \frac{\underline{X}_t}{g^{-1}(t)} = 1$ in probability, where g^{-1} is the inverse function of $g(x) := \mathbb{E}_\infty T_x^-$.
- If X is subcritical, we can show that $\lim_{t \rightarrow 0+} \frac{X_t}{X_t} = 1$, and we have $\frac{X_t}{g^{-1}(t)} = 1$ in probability as $t \rightarrow 0+$.

- Recall that $T_b^- = \int_0^{\tau_b^-} 1/R(Z_s) ds$.
- A **local time** for the process Z is its occupation density.
- Write $l(a, t)$ for the local time of Z at level a up to time t .
- Given $\psi'(0+) \geq 0$, by B. Li and Z. (2017) we have for $x, a > b > 0$,

$$\mathbb{E}_x l(a, \tau_b^-) = W(a - b) - W(a - x)$$

and for $x, a_1, a_2 > b > 0$,

$$\begin{aligned} & \mathbb{E}_x [l(a_1, \tau_b^-) l(a_2, \tau_b^-)] \\ &= [W(a_2 - b) - W(a_2 - x)][W(a_1 - b) - W(a_1 - a_2)] \\ & \quad + W(a_2 - b)[W(a_1 - b) - W(a_1 - x)]. \end{aligned}$$

The moments of T_b^-

Suppose that $\psi'(0+) \geq 0$. Then for $x > b > 0$,

$$\begin{aligned}\mathbb{E}_x(T_b^-) &= \mathbb{E}_x \int_0^\infty R^{-1}(X_s) 1_{s < \tau_b^-} ds \\ &= \int_b^\infty R^{-1}(y) \mathbb{E}_x l(y, \tau_b^-) dy \\ &= \int_b^\infty R^{-1}(y) [W(y-b) - W(y-x)] dy.\end{aligned}$$

$$\mathbb{E}_\infty(T_b^-) = \int_b^\infty R^{-1}(y) W(y-b) dy.$$

X comes down from infinity if and only if

$\int_b^\infty R^{-1}(y) W(y-b) dy < \infty$ for b large enough.

$$\begin{aligned}\mathbb{E}_x(T_b^{-2}) &= \mathbb{E}_x \left(\int_0^\infty R^{-1}(X_s) 1_{s < \tau_b^-} ds \right)^2 \\ &= \int_b^\infty R^{-1}(a_1) da_1 \int_b^\infty R^{-1}(a_2) \mathbb{E}_x l(a_1, \tau_b^-) l(a_2, \tau_b^-) da_2 \\ &= \int_b^\infty R^{-1}(a_1) da_1 \int_b^\infty R^{-1}(a_2) \\ &\quad \times [W(a_2 - b) - W(a_2 - x)][W(a_1 - b) - W(a_1 - a_2)] da_2.\end{aligned}$$

$$\begin{aligned}\mathbb{E}_\infty(T_b^{-2}) &= \int_b^\infty R^{-1}(a_1) da_1 \int_b^\infty R^{-1}(a_2) \\ &\quad \times W(a_2 - b)[W(a_1 - b) - W(a_1 - a_2)] da_2.\end{aligned}$$

If $R(x) = x^\theta$, then

$$\begin{aligned}\mathbb{E}_\infty(T_b^-) &= \int_b^\infty \frac{W(y-b)}{y^\theta} dy = \int_0^\infty \frac{W(y)}{(b+y)^\theta} dy \\ &= \frac{1}{\Gamma(\theta)} \int_0^\infty W(y) dy \int_0^\infty e^{-\lambda(b+y)} \lambda^{\theta-1} d\lambda \\ &= \frac{1}{\Gamma(\theta)} \int_0^\infty e^{-\lambda b} \lambda^{\theta-1} \int_0^\infty W(y) e^{-\lambda y} dy d\lambda \\ &= \frac{1}{\Gamma(\theta)} \int_0^\infty \frac{e^{-\lambda b} \lambda^{\theta-1}}{\psi(\lambda)} d\lambda.\end{aligned}$$

If $R(x) = e^{\theta x}$, then $\mathbb{E}_\infty(T_b^-) = \frac{1}{e^{\theta b} \psi(\theta)}$.

An asymptotic result in the slow regime

Proposition

Suppose that X is subcritical and there exist $0 < c_1 < c_2$ and $\theta > 1$ such that

$$c_1 x^\theta \leq R(x) \leq c_2 x^\theta$$

for all x large enough. Then as $b \rightarrow \infty$,

$$\frac{T_b^-}{\mathbb{E}_\infty T_b^-} \rightarrow 1 \quad \text{in probability under } \mathbb{P}_\infty.$$

Outline of the proof: Check that $\text{Var}_\infty \frac{T_b^-}{\mathbb{E}_\infty T_b^-} \rightarrow 0$.

Under an additional mild condition on ψ we can show the convergence holds almost surely.

Laplace transform of T_b^-

Lemma

Given a locally bounded nonnegative function R on $(0, \infty)$, let $W_n(x)$ satisfy

$$W_0(x) = e^{-\Phi(0)x}, \quad W_{n+1}(x) = \int_x^\infty \frac{W(z-x)}{R(z)} W_n(z) dz, \quad x \geq 0, n \geq 0.$$

Given $b \geq 0$, if $\sum_{n=0}^\infty W_n(b) < \infty$, then for all $x \geq b$,

$$\begin{aligned} \mathbb{E}_x[e^{-\lambda T_b^-}; T_b^- < \infty] &= \mathbb{E}_x \left[\exp \left(-\lambda \int_0^{\tau_b^-} \frac{1}{R(Z_s)} ds \right); \tau_b^- < \infty \right] \\ &= \frac{\sum_{n=0}^\infty \lambda^n W_n(x)}{\sum_{n=0}^\infty \lambda^n W_n(b)}. \end{aligned}$$

Proposition

Suppose $R(x) = x^{\theta_1} e^{\theta_2 x}$, $\theta_2 > 0$, $\theta_1 > 0$, $\psi'(0+) \geq 0$ and

$$\mathbb{E}_{\infty} T_b^- = \int_b^{\infty} \frac{W(y-b)}{R(y)} dy < \infty.$$

Then the sequence $W_n(x)$, $x > b$, can be expressed as $W_0(x) = 1$ and for $n \geq 1$,

$$W_n(x) = \frac{1}{\Gamma(\theta_1)^n} \int_{(0, \infty)^n} \frac{e^{-\sum_{i=1}^n (\lambda_i + \theta_2)x} (\lambda_1 \dots \lambda_n)^{\theta_1 - 1}}{\psi(\lambda_1 + \theta_2) \dots \psi(\sum_{i=1}^n (\lambda_i + \theta_2))} d\lambda_1 \dots d\lambda_n. \quad (2)$$

Let L be a generator on $C^2([0, \infty))$ such that

$$\begin{aligned} Lf(x) &:= x^{\theta_1} e^{\theta_2 x} L^* f(x) \\ &:= x^{\theta_1} e^{\theta_2 x} \left(-bf'(x) + \frac{\sigma^2}{2} f''(x) + \int_0^\infty (f(x+u) - f(x) - uf'(x)) \right) \end{aligned}$$

i.e. L is the generator for the corresponding nonlinear continuous-state branching process with $R(x) = x^{\theta_1} e^{\theta_2 x}$. For $f(x) = e^{-\lambda x}$ we have $L^* f(x) = e^{-\lambda x} \psi(\lambda)$. One can check that for the function W_n given in (2), we have $L^* W_n(x) = x^{-\theta_1} e^{-\theta_2 x} W_{n-1}(x)$. It follows that

$$L \left(\sum_{i=0}^{\infty} \lambda^i W_i \right) (x) = \lambda \sum_{i=0}^{\infty} \lambda^i W_i(x),$$

i.e. $\sum_{i=0}^{\infty} \lambda^i W_i$ is a λ -invariant function of L and it is a “scale function” for process X .

Corollary

Suppose $R(x) = x^\theta$, $\theta > 0$, $\psi'(0+) \geq 0$ and

$$\int_{0+} \frac{\lambda^{\theta-1}}{\psi(\lambda)} d\lambda < \infty.$$

Then the sequence $W_n(x)$, $x > b$ is given by $W_0(x) = 1$ and

$$W_n(x) = \frac{1}{\Gamma(\theta)^n} \int_{(0, \infty)^n} \frac{\exp(-\sum_{i=1}^n \lambda_i x) (\lambda_1 \dots \lambda_n)^{\theta-1}}{\psi(\lambda_1) \dots \psi(\sum_{i=1}^n \lambda_i)} d\lambda_1 \dots d\lambda_n,$$

In particular, for $\psi(\lambda) = \lambda^\alpha$, $1 < \alpha \leq 2$ and $\theta > \alpha$ we have

$$W_0(x) = 1, \quad W_n(x) = x^{n\alpha - n\theta} \prod_{i=1}^n \frac{\Gamma(i\theta - i\alpha)}{\Gamma(i\theta - (i-1)\alpha)}.$$

Corollary

Suppose that $R(x) = e^{\theta x}$, $\theta > 0$ and $\psi'(0+) \geq 0$. Then the sequence $W_n(x)$, $x > b$ is given by $W_0(x) = 1$ and

$$W_n(x) = \frac{1}{e^{n\theta x} \prod_{j=1}^n \psi(j\theta)}.$$

More convergence results

Using the Laplace transforms, we can show that

- if $R(x) = x^\theta$, $\psi(\lambda) = \lambda^\alpha$ with $1 < \alpha \leq 2$ and $\theta > \alpha$, then the distribution of $T_b^- / \mathbb{E}_\infty T_b^-$ under \mathbb{P}_∞ does not depend on b .

$$\mathbb{E}_\infty e^{-\frac{sT_b^-}{\mathbb{E}_\infty T_b^-}} = \left[\sum_{n=0}^{\infty} \left(\frac{s\Gamma(\theta)}{\Gamma(\theta - \alpha)} \right)^n \prod_{i=1}^n \frac{\Gamma(i\theta - i\alpha)}{\Gamma(i\theta - (i-1)\alpha)} \right]^{-1}.$$

- if there exists $1 < \alpha \leq 2$ such that $\psi(\lambda) \sim \lambda^\alpha$ as $\lambda \rightarrow 0+$, and $R(x) = x^\theta$ for $\theta > \alpha$ and all x large enough, then $T_b^- / \mathbb{E}_\infty T_b^-$ converges in distribution under \mathbb{P}_∞ .
- if $R(x) = e^{\theta x}$ for $\theta > 0$ and $\psi'(0+) \geq 0$, then the distribution of $T_b^- / \mathbb{E}_\infty T_b^-$ under \mathbb{P}_∞ does not depend on b .

$$\mathbb{E}_\infty e^{-\frac{sT_b^-}{\mathbb{E}_\infty T_b^-}} = \left[\sum_{n=0}^{\infty} \frac{(\psi(\theta)s)^n}{\prod_{j=1}^n \psi(j\theta)} \right]^{-1}.$$

Small time behavior of \underline{X}_t in the slow regime

Recall that $g(x) := \mathbb{E}_\infty T_x^-$, $x > 0$ and g^{-1} is the inverse function of g .

Lemma

Suppose that process X comes down from infinity and $T_x^-/g(x) \rightarrow 1$ in probability as $x \rightarrow \infty$ under \mathbb{P}_∞ , and for any $h > 1$

$$\liminf_{x \rightarrow \infty} \frac{g(x)}{g(hx)} = c_h > 1.$$

Then we have

$$\frac{\underline{X}_t}{g^{-1}(t)} \rightarrow 1 \text{ in } \mathbb{P}_\infty\text{-probability as } t \rightarrow 0+.$$

Small time behavior of \underline{X}_t in the fast regime

Lemma

Suppose that process X comes down from infinity and $T_x^- / g(x)$ converges in distribution under \mathbb{P}_∞ and for any $h > 1$,

$$\liminf_{x \rightarrow \infty} \frac{g(x)}{g(hx)} = \infty.$$

Then

$$\frac{\underline{X}_t}{g^{-1}(t)} \rightarrow 1 \text{ in } \mathbb{P}_\infty\text{-probability as } t \rightarrow 0+.$$

Two different ways of coming down

Proposition

Suppose that process X comes down from infinity and $\psi'(0+) = \frac{1}{W(\infty)} > 0$. Then \mathbb{P} -a.s., $\lim_{t \rightarrow 0+} \frac{X_t}{\underline{X}_t} = 1$.

Proposition

Suppose that process X comes down from infinity and for $\alpha' > 0$, $W(x) \sim x^{\alpha'}$ as $x \rightarrow \infty$, which by Tauberian theorem is equivalent to $\psi(\lambda) \sim \lambda^{-\alpha'-1}$ as $\lambda \rightarrow 0+$. Then $\limsup_{t \rightarrow 0+} \frac{X_t}{\underline{X}_t} = \infty$.

In summary,

- if X is subcritical and comes down from infinity, then X_t comes down from infinity according to a deterministic function plus relatively small random fluctuations;
- if X comes down from infinity and $W(\infty) < \infty$, then X_t comes down from infinity with large random fluctuations

Given

$$\psi(\lambda) = b\lambda + \frac{\sigma^2\lambda^2}{2} + \lambda^\alpha,$$

- if $R(x) = x^\theta$, then
 - for $b > 0$ and $\theta > 1$,

$$\frac{X_t}{(b(\theta - 1)t)^{1/(1-\theta)}} \rightarrow 1 \text{ in } \mathbb{P}_\infty\text{-probability as } t \rightarrow 0+.$$

- for $b = 0 = \sigma^2$ and $\theta > \alpha$,

$$\frac{X_t}{\left(\frac{\Gamma(\theta)}{\Gamma(\theta-\alpha)}t\right)^{1/(\alpha-\theta)}} \rightarrow 1 \text{ in } \mathbb{P}_\infty\text{-probability as } t \rightarrow 0+.$$

- if $R(x) = e^{\theta x}$, then for $\theta > 0, b > 0$,

$$\frac{X_t}{\theta^{-1} \ln t^{-1}} \rightarrow 1 \text{ in } \mathbb{P}_\infty\text{-probability as } t \rightarrow 0+.$$

Thank you for your attention!